

Making (non)sense of age models: a guide for (ab)users

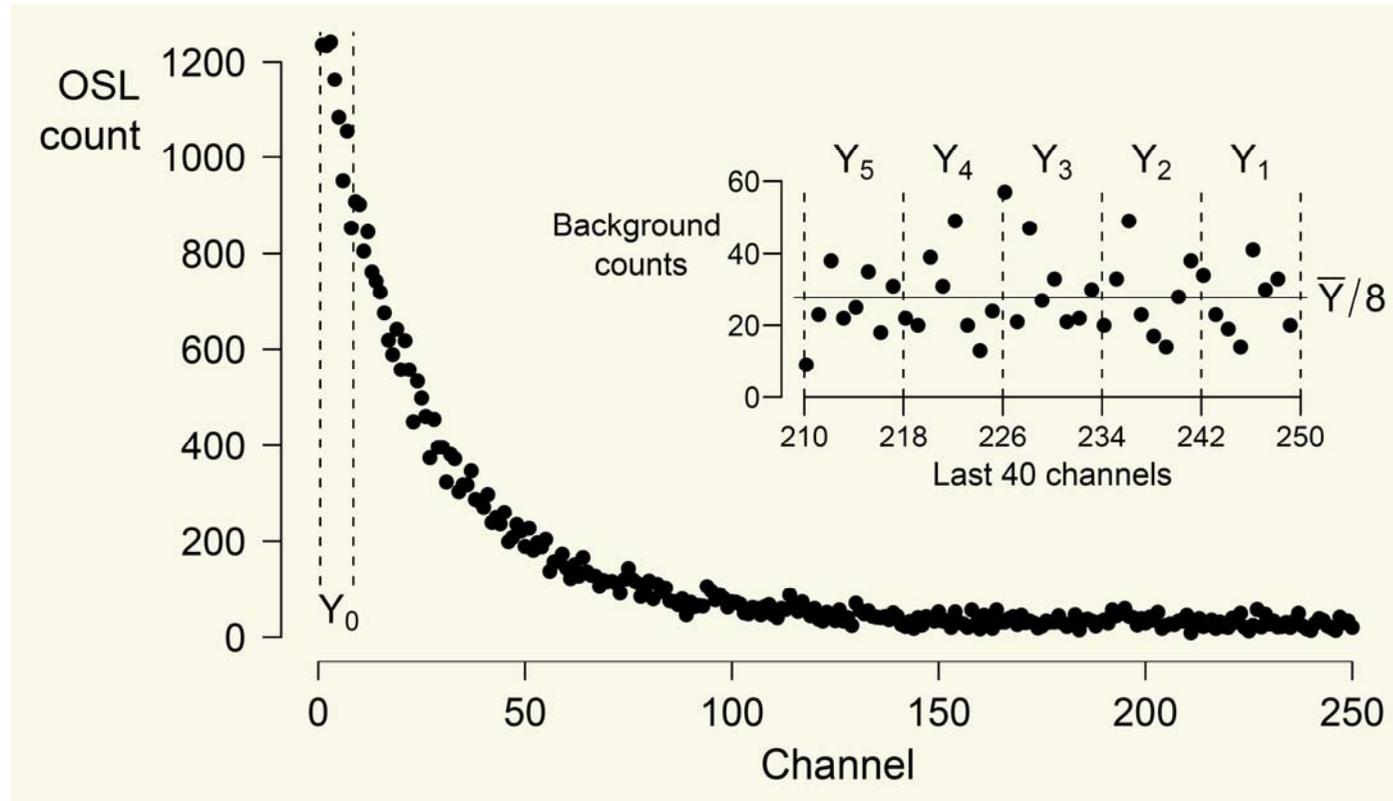
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Talk outline

- 'Age' models for OSL data – why bother?
- Estimating your errors without erring
- Displaying your data simply and correctly
- Data behaving badly – 'over-dispersion'
- Age models – context is everything!
- Age models – some practical tips for beginners

Photon counts from optical stimulation of a quartz grain



OSL counts in $N = 250$ consecutive channels (inset: magnified graph of counts in the last 40 channels). Y_0 = the total count over the first n channels, and Y_1, \dots, Y_k and \bar{Y} are used to estimate the background count. Here, $n = 8$ and $k = 5$.

$$rse(\hat{\mu}_s) = \frac{\sqrt{Y_0 + \bar{Y} / k}}{Y_0 - \bar{Y}}$$

$\hat{\mu}_s$ = estimate of the background-corrected ('net') OSL signal,
assuming the background counts have a Poisson distribution.

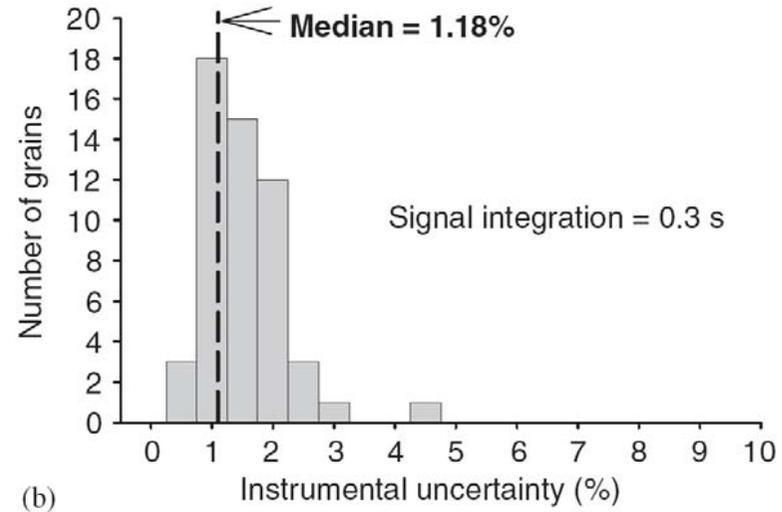
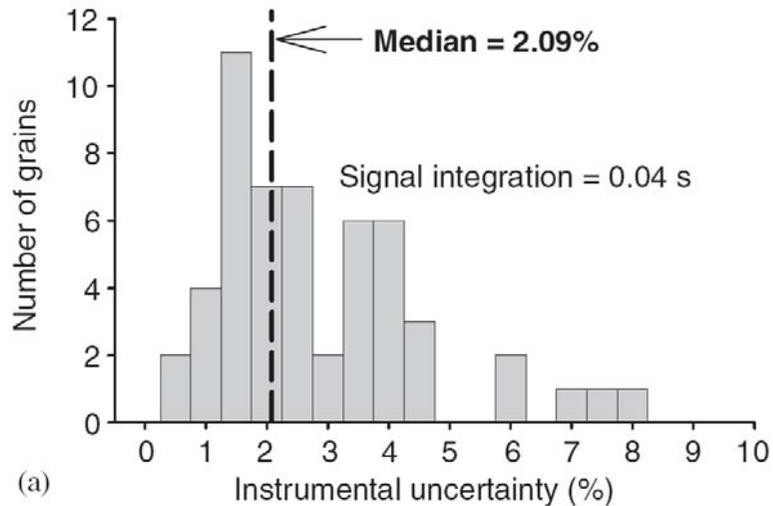
But do they?

- Calculate the ratio of the variance (i.e., the square of the standard deviation) to the mean of the k background counts
- Variance/mean ratio = 1 if Poisson, and >1 if 'extra-Poisson'
- For data with an extra-Poisson component, use equations given by Rex Galbraith (*Ancient TL* 2002):

$$rse(\hat{\mu}_s) \approx \frac{\sqrt{Y_0 + \bar{Y} / k}}{Y_0 - \bar{Y}} \times \sqrt{1 + \frac{\hat{\sigma}^2}{\bar{Y}}} \quad \text{where} \quad \hat{\sigma}^2 = \left(\frac{1}{k-1} \sum_{j=1}^k (Y_j - \bar{Y})^2 \right) - \bar{Y}$$

Other errors to include

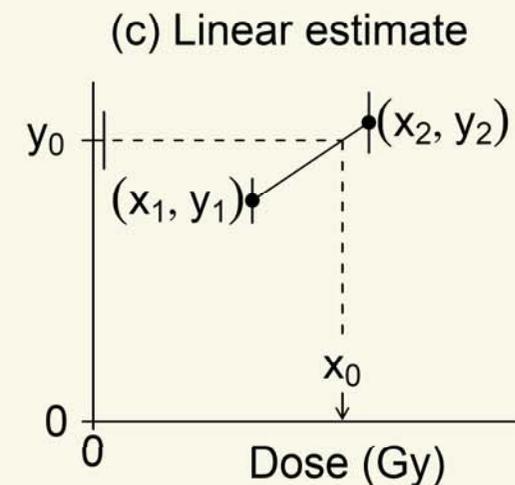
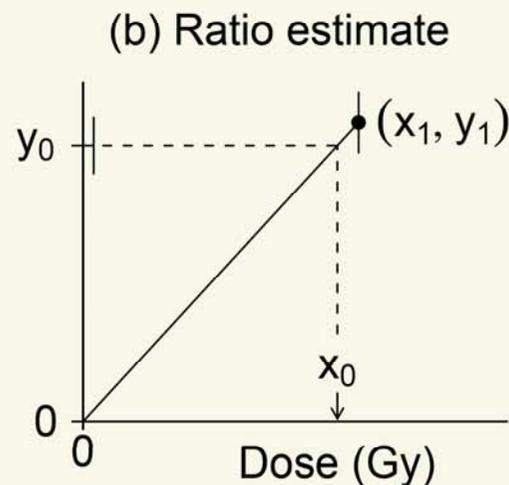
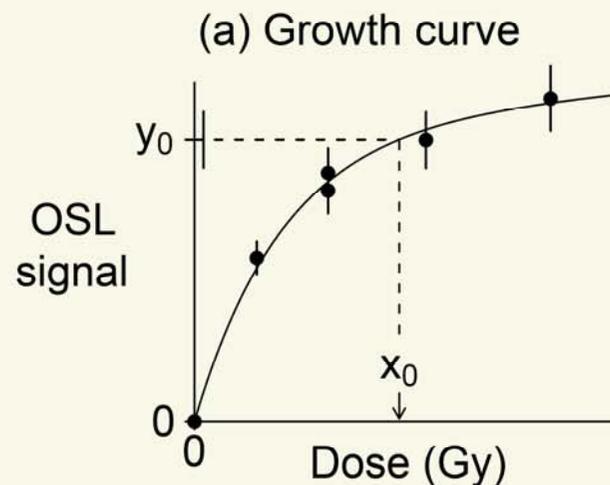
- Instrumental (ir)reproducibility
 - Measure for each instrument
 - Determine for specific signal-integration periods
 - Simple method described by Jacobs *et al.* (*Radiat Meas* 2006)



Estimation of the equivalent dose, D_e

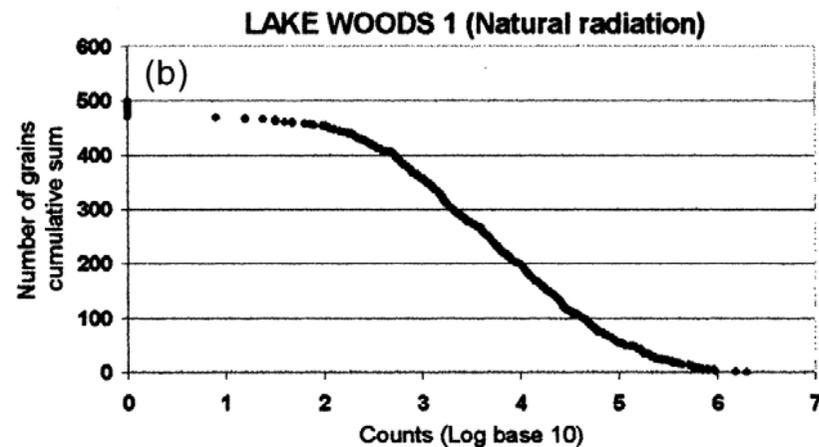
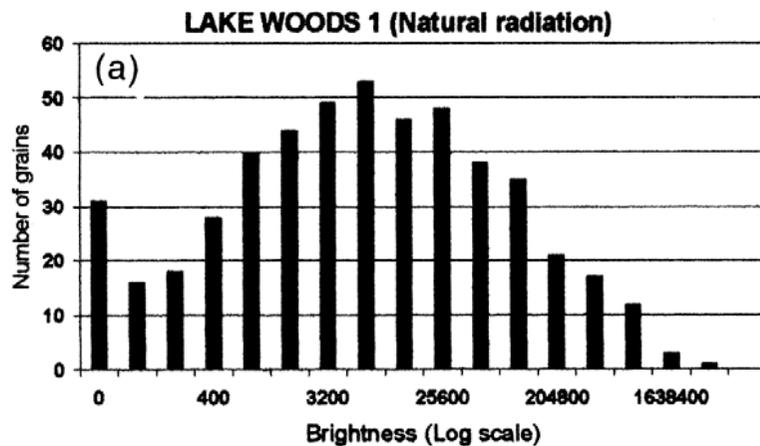
The data are OSL signals y_1, y_2, \dots, y_n (and their precisions) obtained from given known doses x_1, x_2, \dots, x_n (corrected for background and sensitivity change) plus the natural OSL signal y_0 .

We want to estimate the natural dose x_0 that produced y_0 (and its precision).



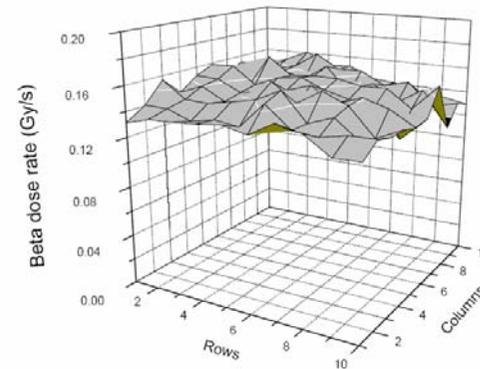
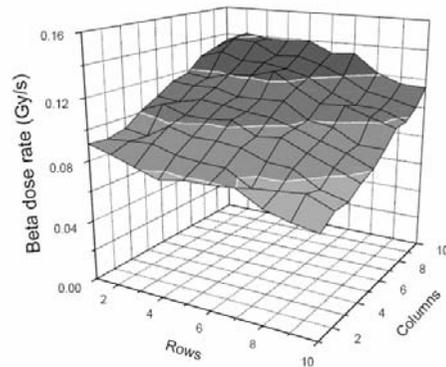
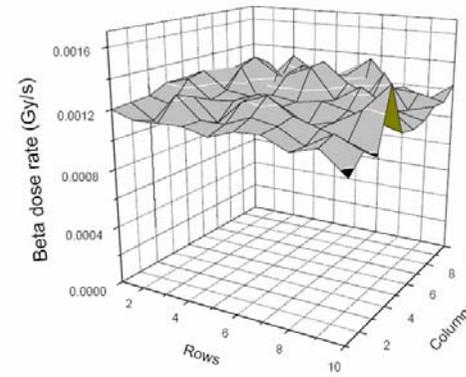
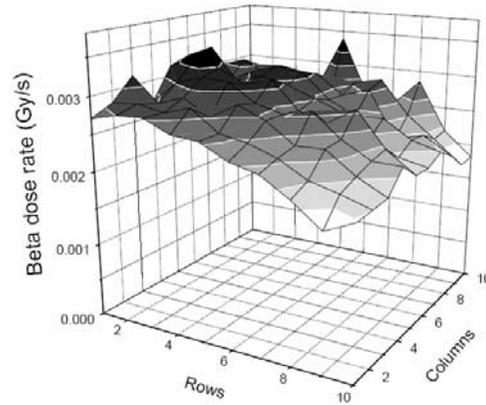
Curve-fitting errors

- Determine by Monte Carlo simulation
 - What form is the frequency distribution of OSL intensities at each dose point – normal (Gaussian) or log-normal?
 - Log-normal distributions of single-grain OSL intensities reported by McCoy *et al.* (*Radiat Meas* 2000)
 - Appropriate MC methods described by Yoshida *et al.* (*Radiat Meas* 2003) and Duller (*Ancient TL* 2007)



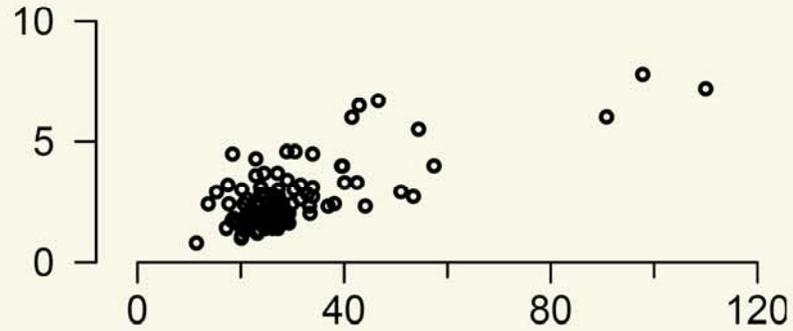
Beta-source variability to single grains

- Cross-disc spatial variation in dose rate from laboratory β sources
- Measure for each instrument, e.g. Ballarini *et al.* (*Ancient TL 2006*)

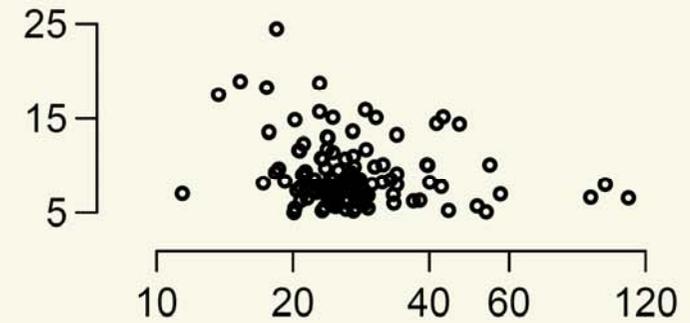


D_e estimates for 120 aliquots

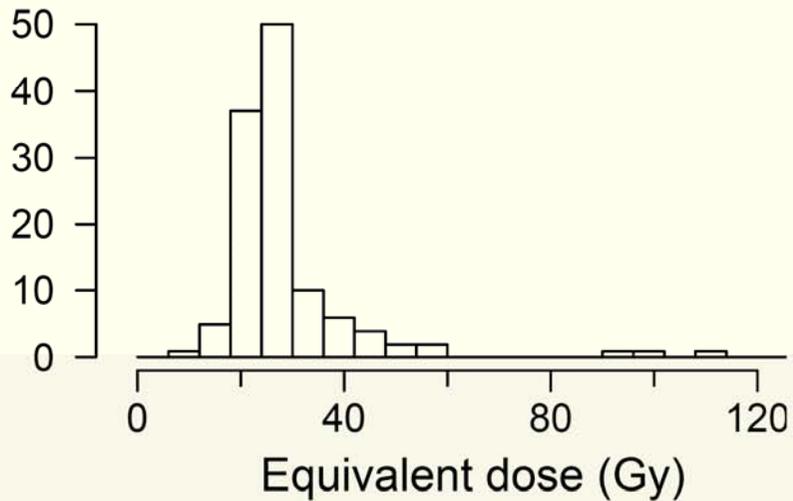
Standard error (Gy)



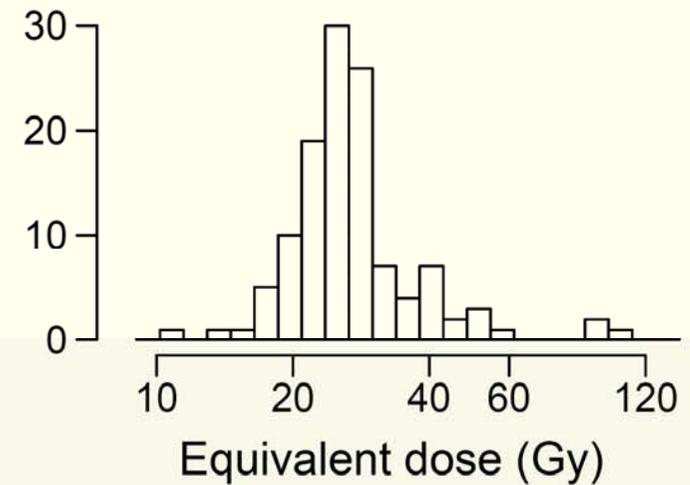
Relative standard error (%)



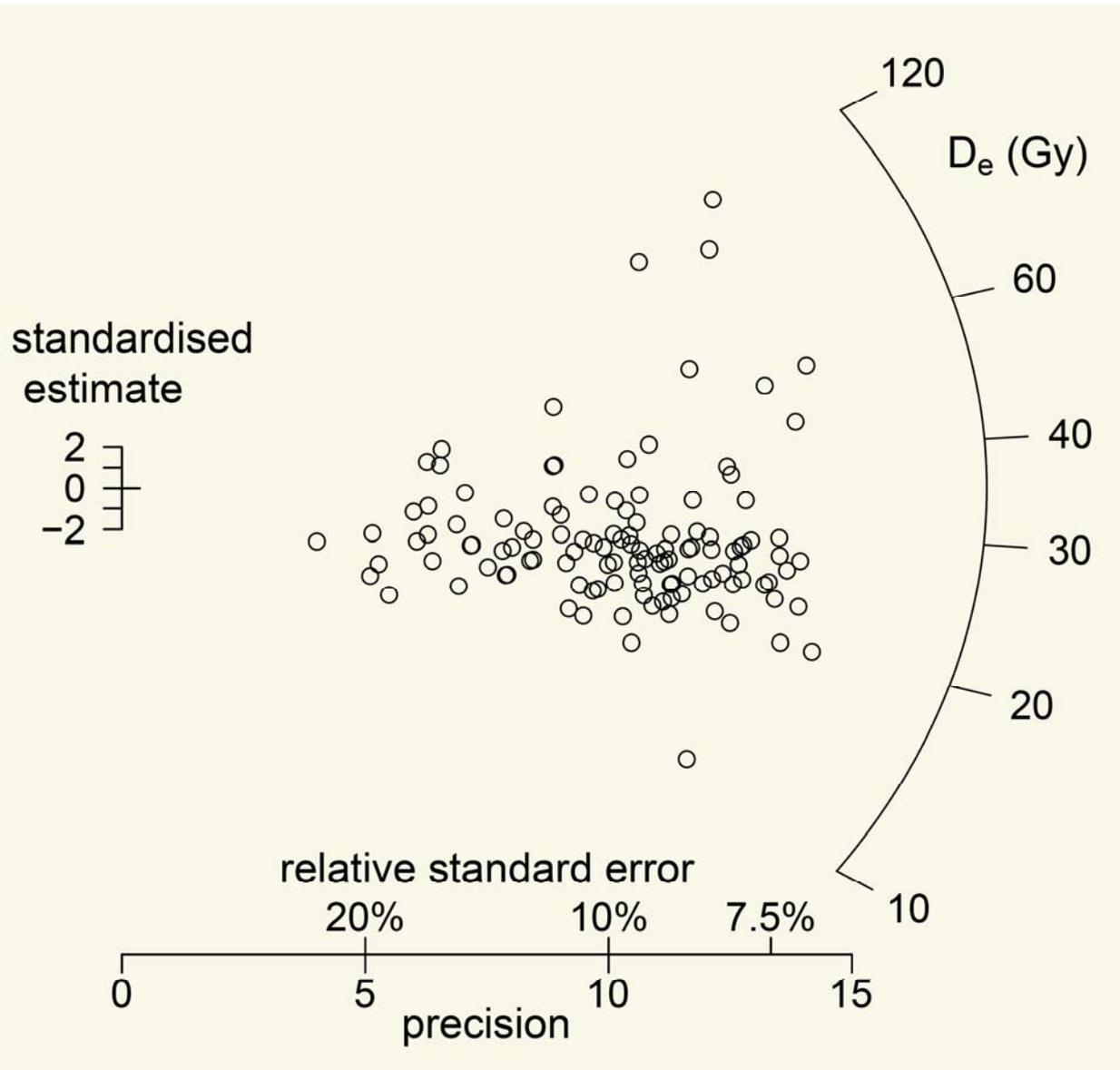
Number of aliquots



Number of aliquots



Radial plot of D_e estimates for 120 aliquots



Radial plots

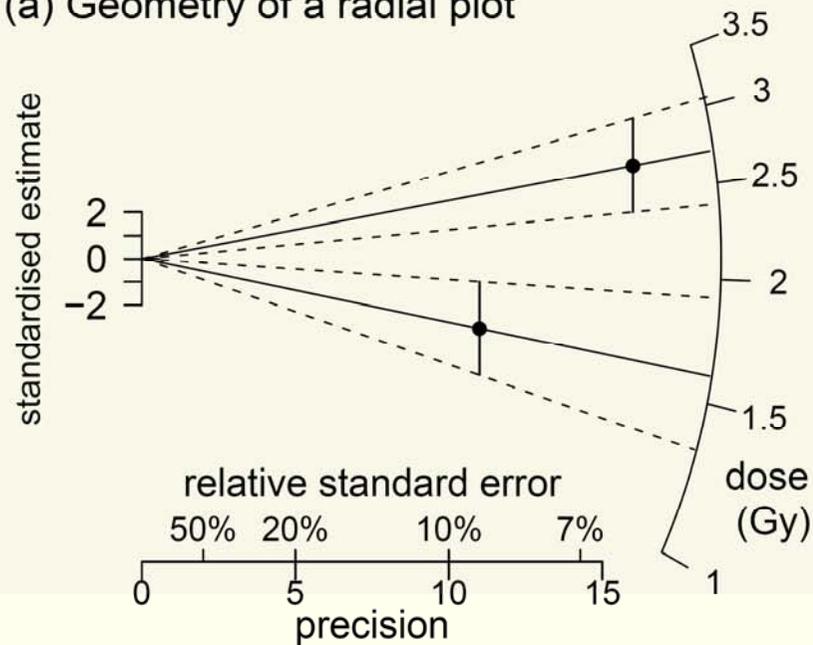
Data are estimates z_1, z_2, \dots, z_n with standard errors $\sigma_1, \sigma_2, \dots, \sigma_n$

Draw a scatter plot of $y_i = (z_i - z_0)/\sigma_i$ against $x_i = 1/\sigma_i$

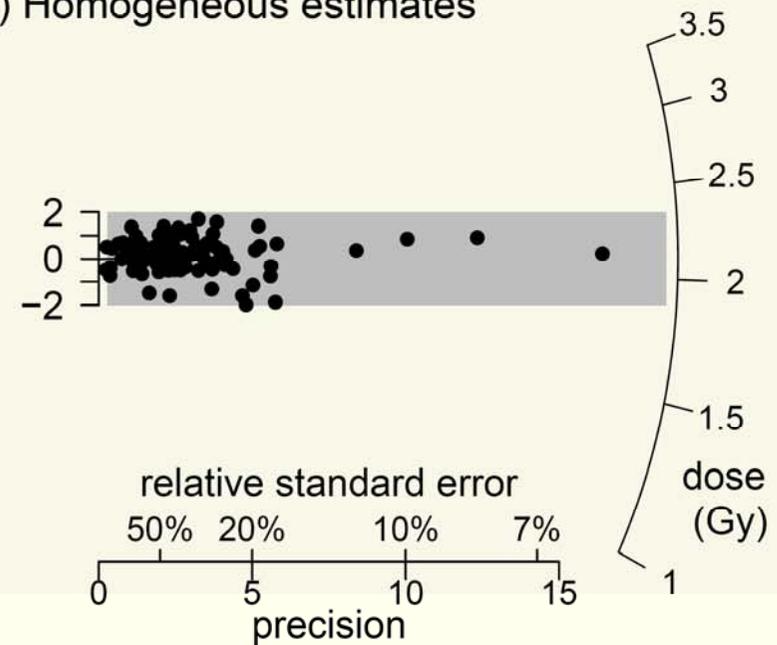
Then $z_i - z_0 = y_i/x_i = \text{slope of line through } (0,0) \text{ and } (x_i, y_i)$

and each y_i has unit standard deviation

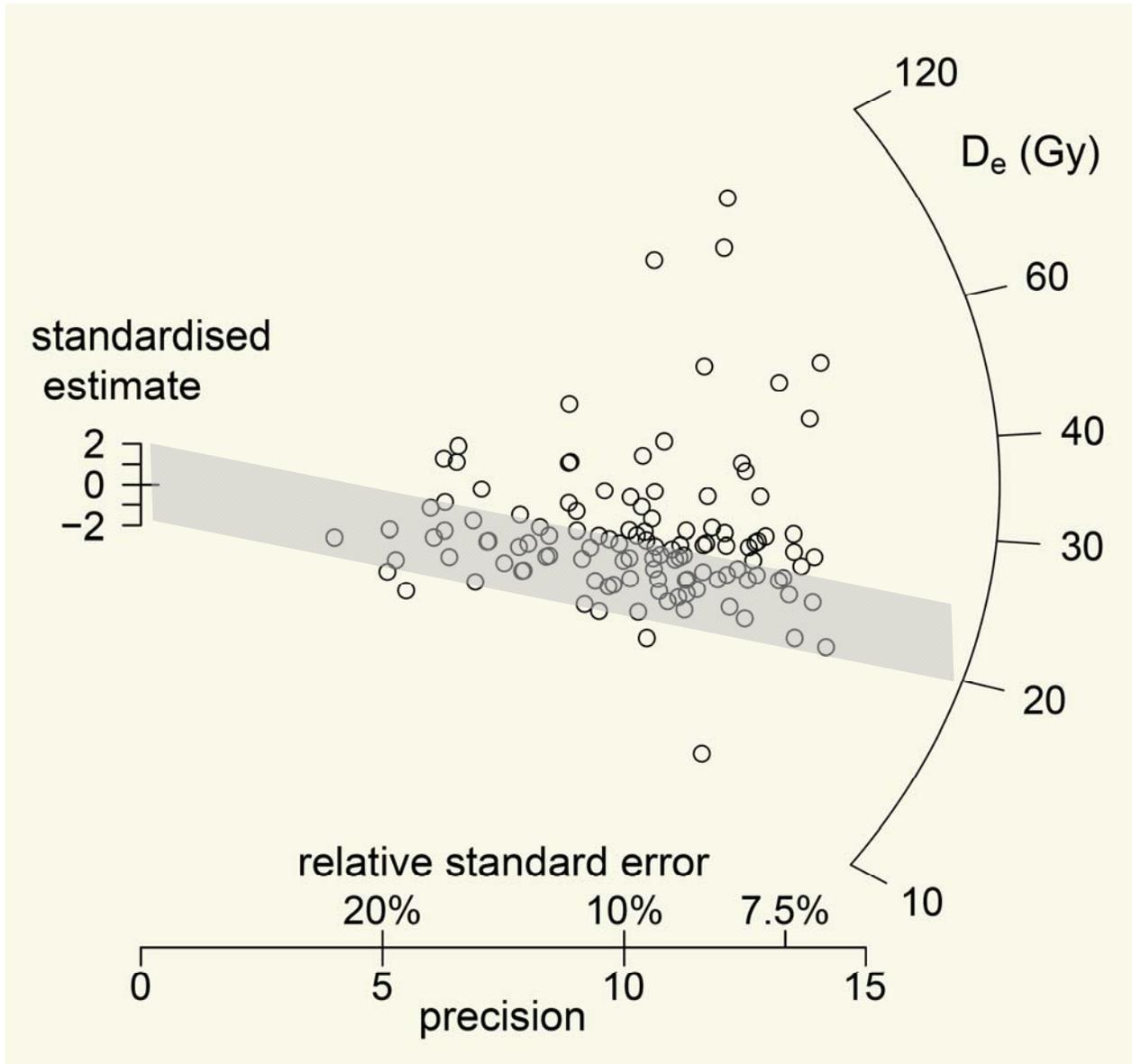
(a) Geometry of a radial plot



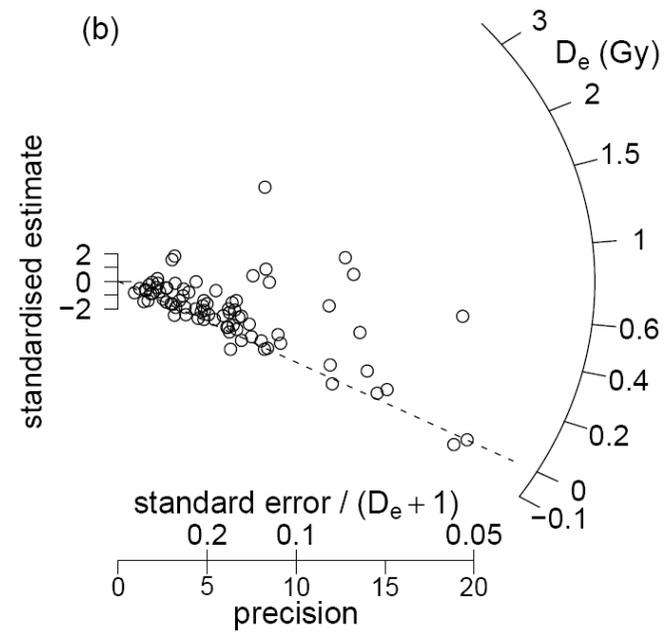
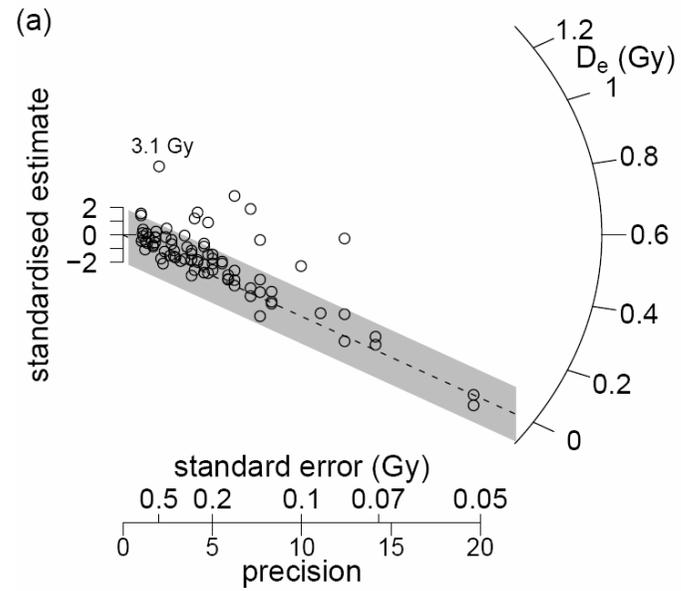
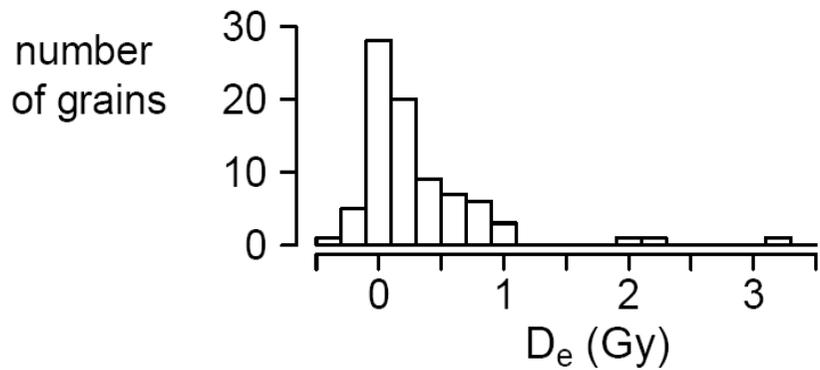
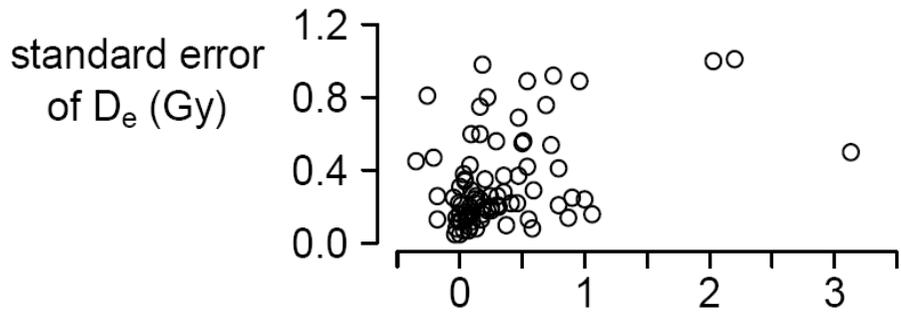
(b) Homogeneous estimates



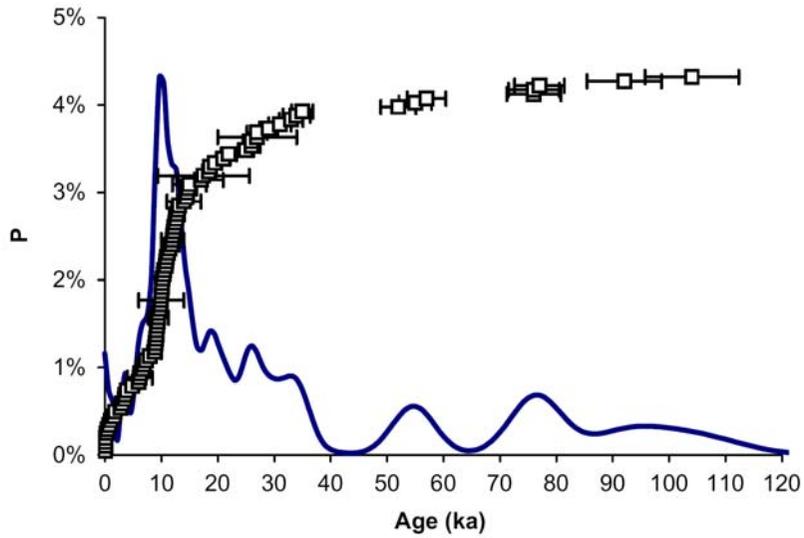
Radial plot of D_e estimates for 120 aliquots



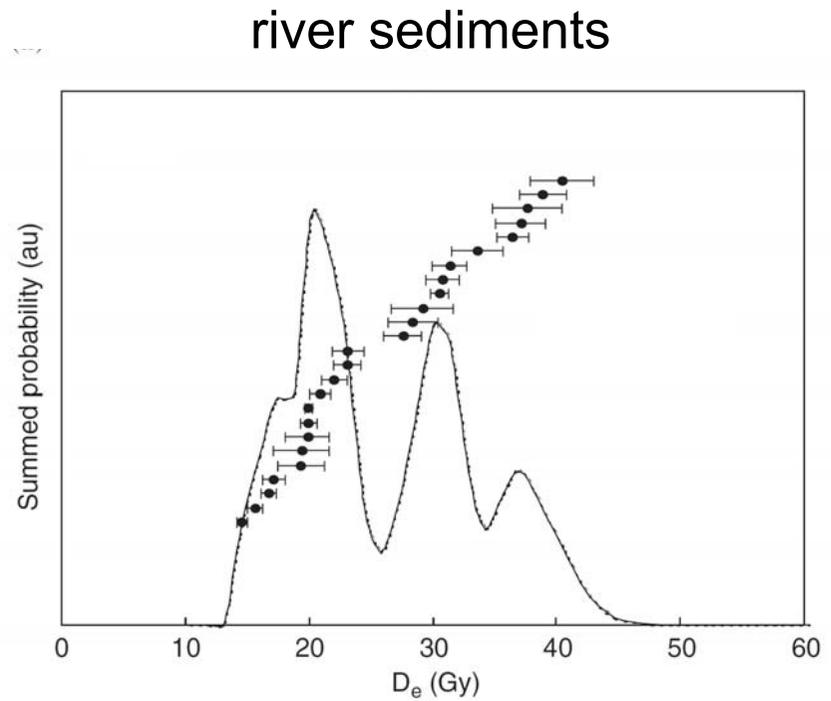
D_e estimates for young and modern samples



Weighted histograms (probability density plots)

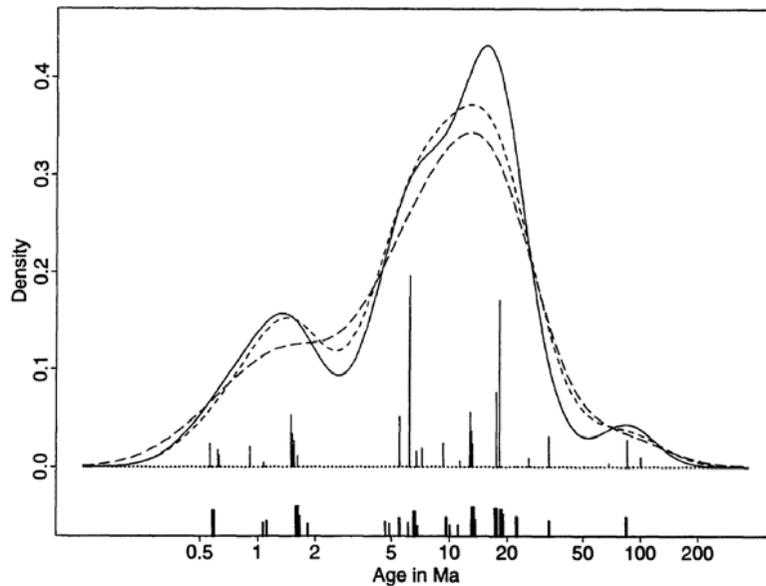
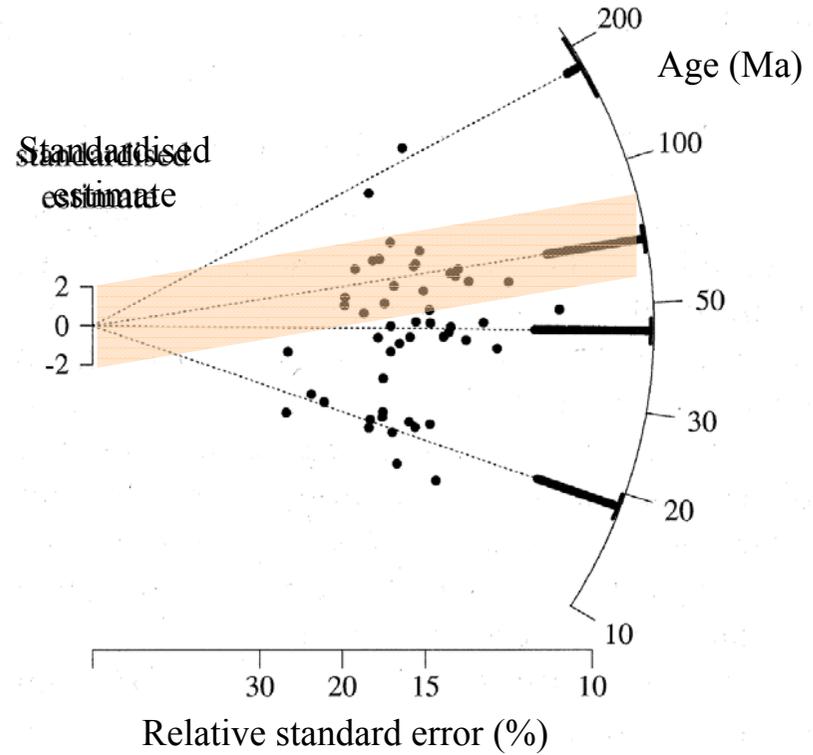
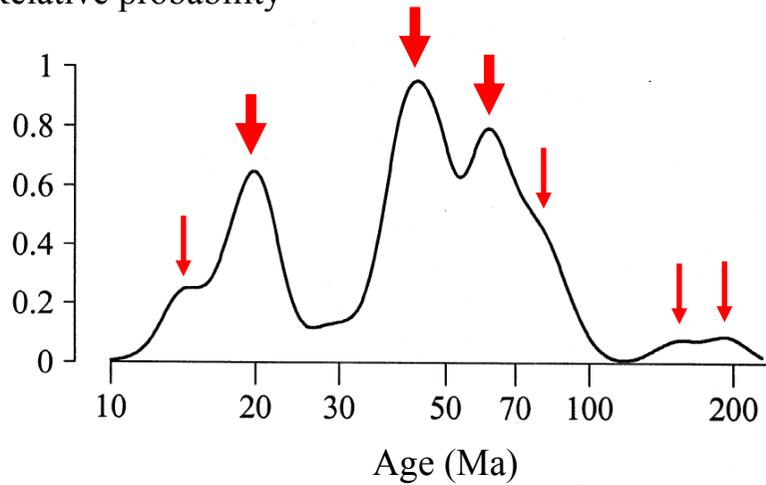


dune sediments



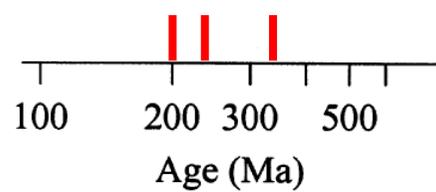
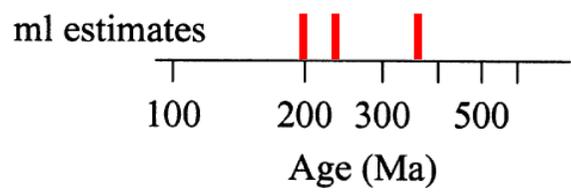
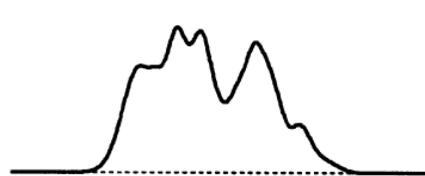
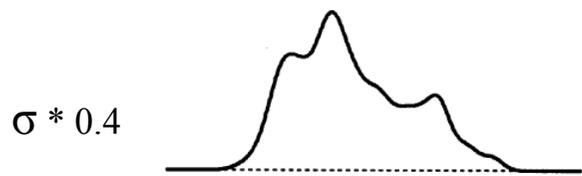
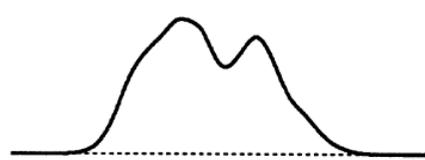
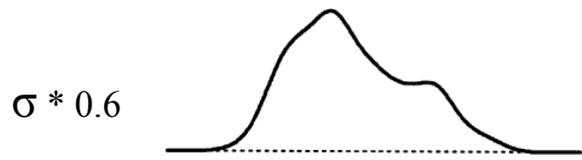
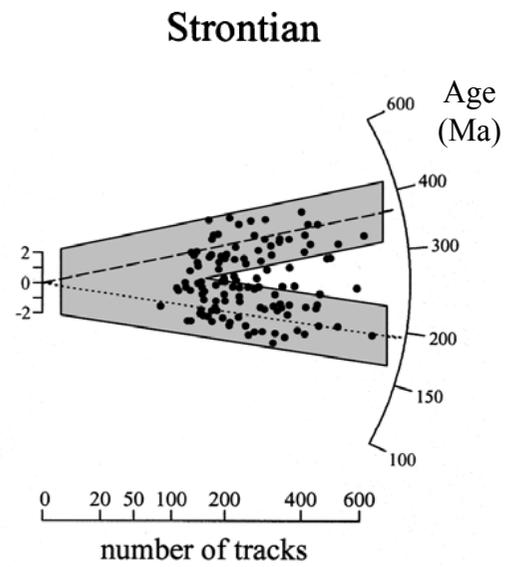
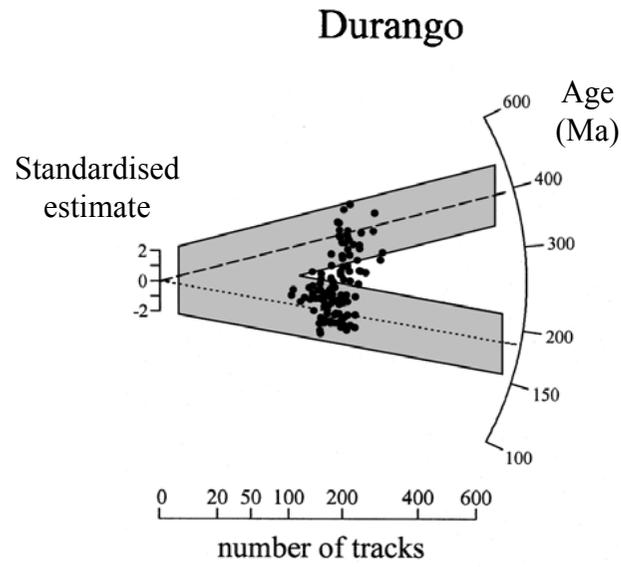
Probability density plots: conflating 2 sources of error

Relative probability



Galbraith (*Radiat Meas* 1998)

Goutis (*J Am Stat Assoc* 1997)



Over-dispersion

- The extra spread in D_e values remaining after having taken into account measurement uncertainties:
 - photon counting statistics
 - background correction (extra-Poisson?)
 - instrumental (ir)reproducibility
 - curve-fitting errors
 - beta-source heterogeneity (for single grains)
- Relative standard deviation of the D_e estimates above and beyond all known uncertainties, expressed in %
- Denoted by σ_b by Galbraith *et al.* (*Radiat Meas* 2005)
- Ubiquitous in natural samples and lab-dosed samples
 - dose-recovery tests
 - recycling ratios

Why do we get over-dispersion?

- Because of 'experimental' and 'natural' variation
 - experimental errors are reducible
 - but natural variation is inherent
- Variation between aliquots/grains in the lab
 - OSL source traps not filled to same extent
 - before burial: partial bleaching
 - after burial: sediment mixing, micro-dosimetry variations
 - heating/bleaching don't empty all traps uniformly
 - differences in thermal transfer effects
- Non-identical field and lab conditions
 - bleaching spectra
 - types of ionising radiation
 - dose rates
 - defect migration with time/temperature

How much over-dispersion do we get?

- Factorial experiment (Galbraith *et al.*, *Radiat Meas* 2005)
- Six replicates of a 2^3 factorial design. The 3 factors:
 1. size of test dose – 0.5 or 5 Gy
 2. preheat temperature – 180°C or 260°C
 3. size of aliquot – ‘small’ (8 grains) or ‘large’ (80 grains)

S1	K166, annealed then given a laboratory gamma dose of 2.74 Gy
S2	K162, bleached then given a laboratory beta dose of 46 Gy
S3	AC150, bleached then given a laboratory beta dose of 25 Gy
S4	AC150, natural (field) dose of ~ 25 Gy
S5	K162, bleached then given a laboratory beta dose of 2.74 Gy
S6	K162, natural (field) dose of ~ 46 Gy

Sample and dose	Aliquot size	Recycling ratio				Natural and surrogate natural			
		σ_b	(95% CI)	σ_w	σ_e	σ_b	(95% CI)	σ_w	σ_e
S1: K166, γ dose 2.74 Gy	Small	1.3	(0.8–2.0)	2.9	3.2	4.7	(3.5–6.5)	2.7	5.4
	Large	0.9	(0.6–1.3)	0.4	1.0	0.8	(0.6–1.2)	0.4	0.9
S2: K162, β dose 46 Gy	Small	1.5	(0.9–2.5)	2.5	2.9	1.3	(0.5–2.5)	2.4	2.7
	Large	1.0	(0.6–1.8)	2.4	2.6	1.5	(0.9–2.5)	2.4	2.8
S3: AC150, β dose 25 Gy	Small	0.7	(0.0–1.8)	3.3	3.4	1.6	(0.7–3.0)	3.4	3.8
	Large	0.7	(0.3–1.2)	1.2	1.4	1.6	(1.1–2.4)	1.3	2.1
S4: AC150, field dose \sim 25 Gy	Small	0.9	(0.0–2.0)	2.4	2.6	7.1	(5.2–10.0)	2.4	7.5
	Large	0.9	(0.5–1.6)	1.4	1.6	3.7	(2.8–5.2)	1.3	3.9
S5: K162, β dose, 2.74 Gy	Small	0.8	(0.0–2.9)	5.5	5.6	11.5	(8.2–16.6)	5.1	12.6
	Large	0.8	(0.0–1.5)	1.4	1.6	3.1	(2.3–4.4)	1.4	3.4
S6: K162, field dose \sim 46 Gy	Small	2.6	(1.7–4.2)	3.7	4.5	17.7	(13.6–24.3)	3.3	18.0
	Large	2.5	(1.8–3.5)	1.9	3.1	11.8	(9.0–16.3)	1.7	11.9
Combined estimate of σ_b for R_1 vs. R_2		1.4	(1.2–1.6)						

Estimates of over-dispersion

- Up to 3% for recycling ratios
- Several % for dose recovery tests (surrogate naturals)
- More for natural samples (typically 10–20%)

Age models (actually, D_e models)

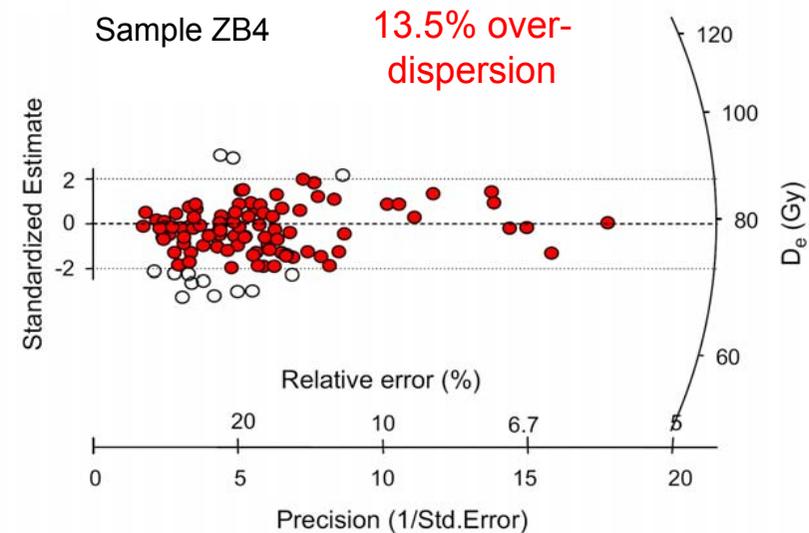
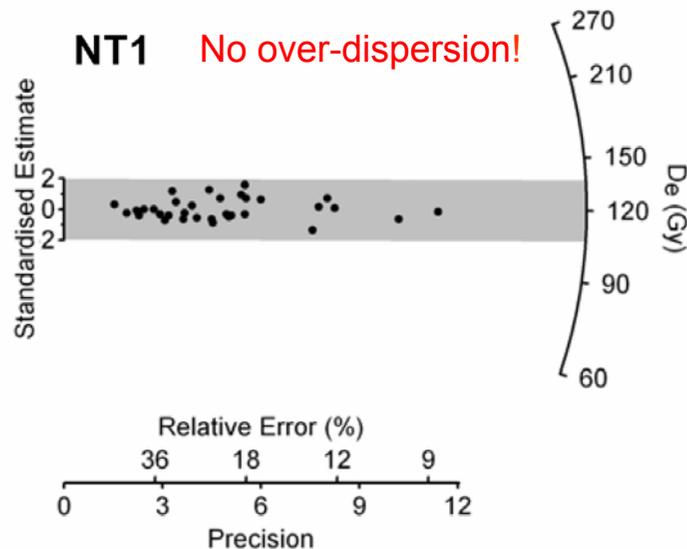
- Typically, the weighted mean of $Y_1, Y_2 \dots$ etc. is calculated as:

$$\bar{Y} = \frac{Y_1 / \sigma_1^2 + Y_2 / \sigma_2^2 + \dots}{1 / \sigma_1^2 + 1 / \sigma_2^2 + \dots}$$

- Note that the weights are by absolute (not relative) standard error
- So the weighted mean of D_e values with the same relative errors will be biased towards the smaller values
- Now, the relative standard error of an estimate \approx absolute standard error of its natural logarithm (\log_e)
- So, by using the natural logs of the D_e values, the weighted mean can be calculated with the relative standard errors
- The 'Central' and 'Common' age models do this

Central and Common age models

- Developed for fission-track ages by Galbraith & Laslett (*NTRM* 1993)
- Extended to D_e estimates by Galbraith *et al.* (*Archaeometry* 1999)
- Both give weighted mean (\approx geometric mean of true D_e values)
- Central age model calculates and includes any over-dispersion



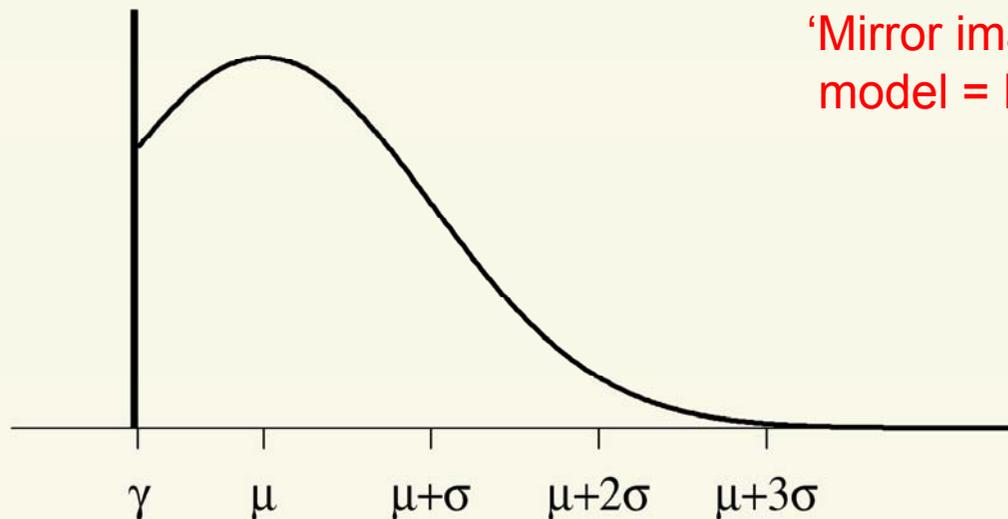
When (not) to use these models

- Central/Common age models may be appropriate for sediments that:
 - were well bleached when deposited
 - have remained undisturbed since burial
 - and have no dosimetry complications ('hot' spots, 'cold' spots)
- Probably not suitable for poorly bleached and/or mixed samples
- Alternative models include:
 - Minimum age model (Galbraith & Laslett, Galbraith *et al.*)
 - Maximum age model (Olley *et al.*, *Quat Sci Rev* 2006)
 - Finite mixture model (Galbraith & Green, *NTRM* 1990; Roberts *et al.*, *Radiat Meas* 2000)
- Key points:
 - field context is vital
 - take all available evidence into account to decide model choice
 - obtain independent estimate of over-dispersion for MAM & FMM
 - don't apply every model to every sample!

Continuous mixtures

- Where the smallest D_e values \approx age of 'target' event
- 3- and 4-parameter 'Minimum' age models:

Here x_i has a mixture distribution — a mixture of a single minimum value γ with probability p and a continuous range of values greater than γ , usually modelled as a truncated normal distribution.

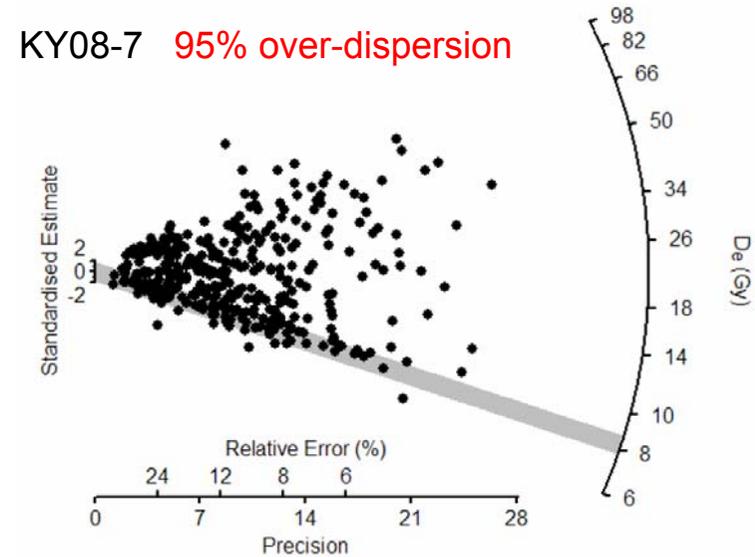


'Mirror image' of Minimum age model = Maximum age model

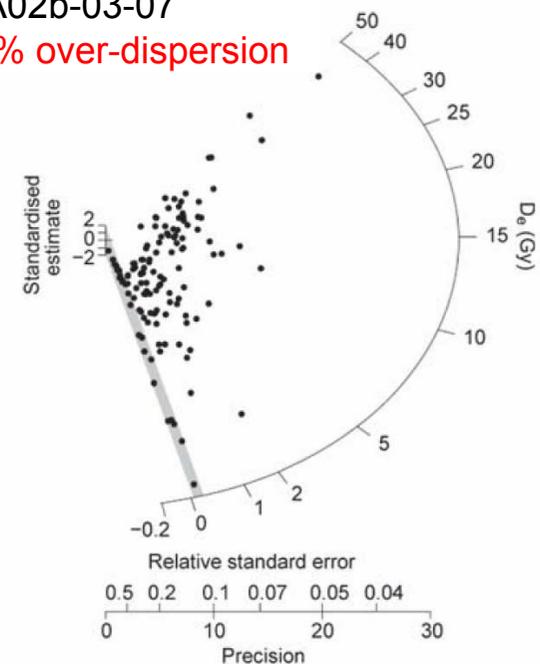
There are 4 parameters: γ , μ , σ and p , the proportion of grains that have the minimum dose. $\gamma = \mu$ in 3-parameter version.

Minimum age model

- Add an estimate of over-dispersion to each measurement error (in quadrature) before running model
- Standard (logged) version of model appropriate for most natural samples
- Modern or young samples, use un-logged version (Arnold *et al.*, *Quat Geochron* 2009)
- Check fits using formal statistical criteria (e.g. maximum log likelihood)

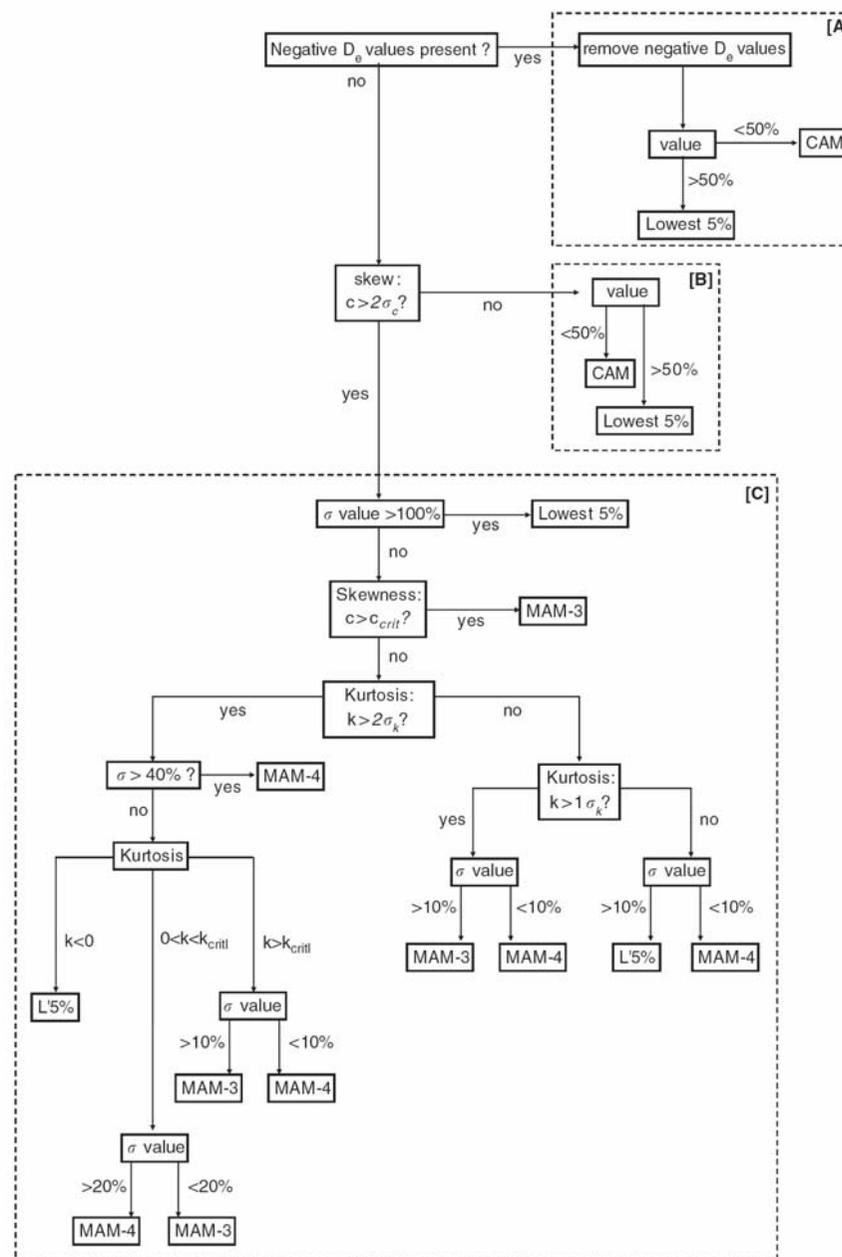


USA02b-03-07
119% over-dispersion



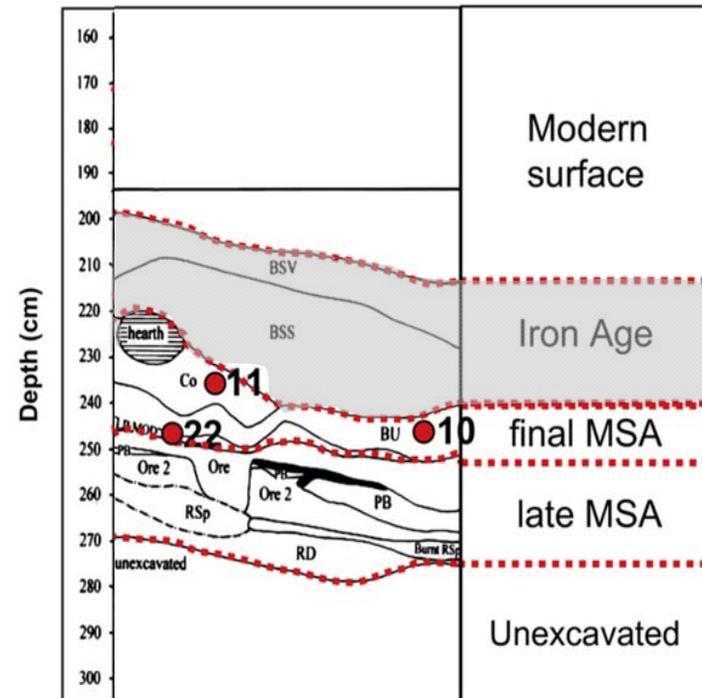
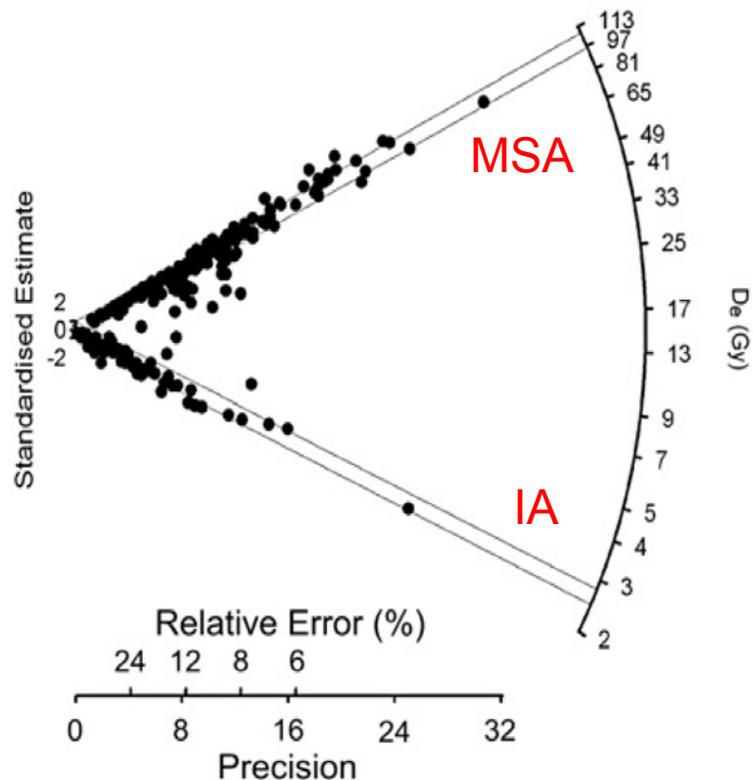
Flowcharts – friend or foe?

- Model decision based on simulated data (Bailey & Arnold, *Quat Sci Rev* 2006)
- Validity depends on what data went into model
- ‘Real world’ complications, e.g. between-sample variation in over-dispersion
- Skewness and kurtosis:
 - conflate 2 sources of error
 - should be logged values
- Don't rely on flowcharts



Discontinuous mixtures

- 'Finite' mixture model for samples composed of discrete age-populations of grains mixed together after burial
 - common at archaeological sites (Jacobs & Roberts, *Evol Anthrop* 2007)
 - e.g. sample 11 from Sibudu Cave (South Africa) is a mixture of Iron Age (IA) and Middle Stone Age (MSA) sediments



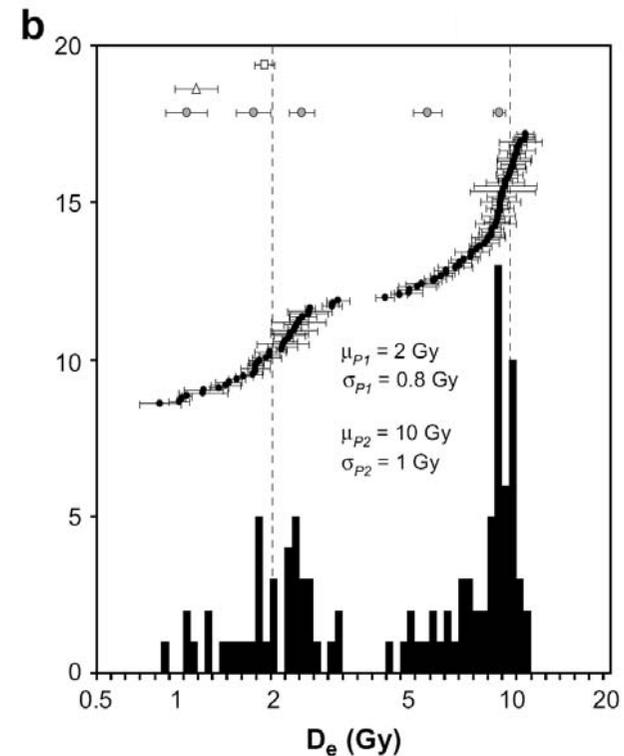
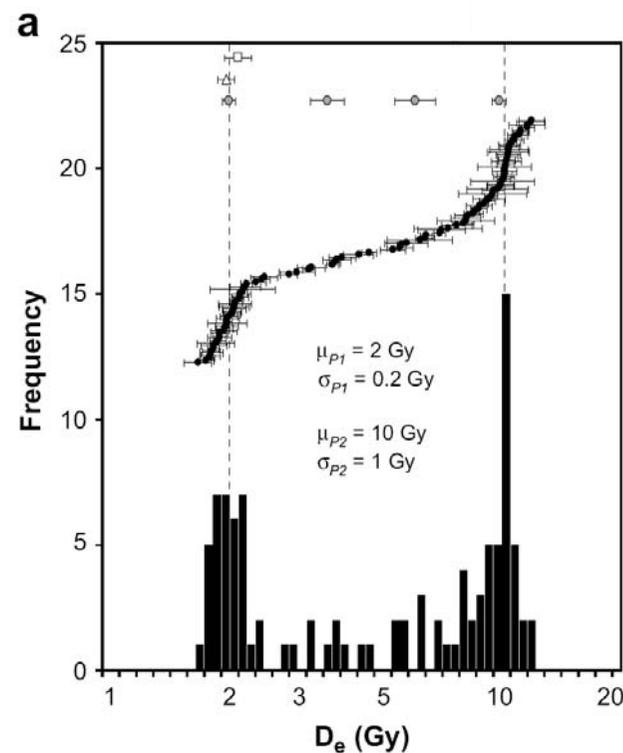
Finite mixture model

- Apply only to single grains and not to continuous mixtures!
- **Input:**
 - single-grain D_e values and standard errors
 - estimated number of finite D_e components
 - estimate of inherent over-dispersion (σ_b)
- **Output:**
 - D_e and standard error of each component (Central age model)
 - relative proportion of grains in each component
 - two estimates of goodness-of-fit
 - maximum log likelihood (MLL)
 - Bayes Information Criterion (BIC)
- FMM developed for fission-track ages by Galbraith & Green (1990)
- But fission-track ages have zero over-dispersion, so OSL version (Roberts *et al.*, 2000) includes over-dispersion as extra parameter

- **Procedure:**

- **change number of components and over-dispersions to:**
 - maximise MLL (increase by > 2 for each added component)
 - minimise BIC (MLL penalised for each added component)
- **MLL and BIC may not always give same best-fit outcomes, so check result not sensitive to different over-dispersion values**
 - David *et al.* (*J Quat Sci* 2007), Jacobs *et al.* (*J Arch Sci* 2008)

FMM will detect 'phantom' D_e populations even for very small aliquots, e.g. 4-grain aliquots with only 1–2 luminescent grains (Arnold & Roberts, *Quat Geochron* 2009)



Concluding remarks

- Models are garbage collectors, so collect good data!
- Correctly estimate your measurement errors
- Display your D_e values meaningfully (as radial plots)
- Measure (and report) your D_e over-dispersion values
- Consider sample context when selecting an age model
- Include over-dispersion when implementing age model
- Don't use FMM for m-g aliquots or continuous mixtures